

A Meta-Analysis of Childhood Maltreatment in Relation to Psychopathic Traits

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Abstract

Psychopathy is a personality disorder characterized by a mix of traits belonging to four facets: affective (e.g., callous/lack of empathy), interpersonal (e.g., grandiosity), behavioral instability (e.g., impulsivity, poor behavioral controls), and social deviance (e.g., juvenile delinquency, criminal versatility). Several scholars have argued that early childhood maltreatment impacts the development of psychopathy, although views regarding its role in the four facets differ. We conducted a meta-analysis including 47 studies comprising a total of 349 effect sizes and 12,737 participants, to investigate the association between the four psychopathy facets and four types of child maltreatment: physical abuse, emotional abuse, neglect, and sexual abuse.

We found support for a moderate link between overall psychopathy and childhood physical abuse, emotional abuse, and neglect, as well as overall childhood maltreatment. The link between psychopathy and childhood sexual abuse was small, but still significant. These associations were stronger for the behavioral and antisocial facets than for the affective and interpersonal facets of psychopathy, but nearly all associations were significant. Our findings are consistent with recently developed theories on the role of complex trauma in the development of severe personality disorders. Trauma-focused preventive and therapeutic interventions can provide further tests of the trauma-psychopathy hypothesis.

Keywords: Psychopathy, childhood maltreatment, meta-analysis, complex trauma, treatment

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I had a kind of rough childhood. My Mom was admitted to a psychiatric institution many times, and my Dad was extremely harsh. He and my older brothers used to hit me with a broom. They also made me sit in a corner with hot peppers in my mouth. My memories of my childhood are rather chaotic, I do not like to think about that period. I was bullied in elementary school. But then, one day, I remember I hit this kid in the school yard during recess, and that's when I found out I could be a bully too. And from that time on, I switched from being a victim to being the perpetrator. (Patient A.)

This quote is taken from a biographical interview administered in connection with the coding of the Psychopathy Checklist-Revised (PCL-R) for a 35-year old male violent offender admitted to a secure forensic psychiatric hospital. His criminal history is long and versatile; he assaulted and cheated on multiple intimate partners, embezzled money from a former employer, and he takes pride in his ability to manipulate others. He lacks remorse for any of his wrongdoings. He reports his family home environment was chaotic and he suffered maltreatment as a child, but he does not want to dwell on this. His consensus PCL-R score, based on the interview and extensive collateral information, amounted to 38, with 40 being the maximum score on the instrument.

Psychopathic personality disorder or psychopathy, as currently defined, consists of several dimensions of symptoms (Patrick, 2018; Sellbom & Drislane, 2020). The most prominent among the dimensional models of psychopathy is the one developed by Robert Hare and colleagues, on the basis of his work with the PCL-R and related instruments (e.g., Psychopathy

Checklist: Youth Version; Forth et al., 2003; Youth Psychopathic Traits Inventory; Andershed et al., 2007; Antisocial Process Screening Device; Frick & Hare, 2001). Hare, in turn, was inspired by the work of Cleckley (1941, 1988) who had derived a set of 16 criteria from his clinical work with psychiatric patients (Hare & Neumann, 2008). Hare's PCL-R comprises two higher order factors (Interpersonal/Affective [Factor 1] and Social Deviance [Factor 2]), which exhibit moderate to strong intercorrelations (Hare, 1991), and can be further divided into four facets: Interpersonal and Affective (from Factor 1), Lifestyle and Antisocial (from Factor 2). This 4-facet structure has been repeatedly supported by Confirmatory Factor Analysis and Structural Equation Modeling using different PCL-scales in samples of adolescents and adults, from community, offender and clinical settings, from North America and Europe (e.g., Neumann et al., 2015; for an overview see Hare et al., 2018).

Until a little over a decade ago, empirical research on psychopathy was largely equivalent to research about PCL-derived psychopathy. Since then, several other psychopathy models and measures have been developed and tested, the majority of which are based on self-report. This has resulted in psychopathy research moving into subclinical and so-called 'successful psychopathic' samples (Benning et al., 2018). One example is the Psychopathic Personality Inventory (PPI/PPI-R), which comprises two largely unrelated higher-order factors: Fearless Dominance and Self-Centered Impulsivity (Benning et al., 2003; Lilienfeld & Widows, 2005). Another influential psychopathy conceptualization is Patrick's Triarchic Model of Psychopathy (TriPM; Patrick et al., 2009). The TriPM delineates psychopathy as three partially overlapping, but distinct subdimensions: Boldness (e.g., social dominance, low stress reactivity, and thrill-adventure seeking), Meanness (e.g., callousness, coldheartedness, and antagonism), and Disinhibition (e.g., impulsivity and negative affectivity; Patrick et al., 2012; Sleep et al., 2019).

Boldness is strongly correlated with PPI Fearless Dominance (Anderson & Kiehl, 2014; Sellbom & Phillips, 2013) and purportedly reflects “assertiveness and persuasiveness, bravery, and venturesomeness” (Patrick et al., 2009).

Etiology of Psychopathy: Theories about the Role of Childhood Maltreatment

Given the large societal costs and interpersonal harm caused by individuals with psychopathy, insight into its etiology could help pave a path towards effective prevention and treatment. If it can be demonstrated that environmental factors, such as childhood maltreatment, play a role in the development of psychopathy, these factors could be targeted for intervention. Research on the heritability of psychopathic traits has uncovered that around 50% of the variance in psychopathic traits, and the affective (callous-unemotional) traits in particular, can be accounted for by genetic factors (Fontaine et al., 2010; Viding et al., 2005; for a review see Waldman et al., 2018). However, it needs to be acknowledged that a discovery of heritability of psychopathic traits should not be equated with a genetic or neurobiological etiology because genes, brain, and behavior are in dynamic interaction throughout life, with opportunities for both improvement and deterioration of mental health problems (Wermter et al., 2010).

Speculating about the etiology of psychopathy, Cleckley (1941, 1988) tended more towards an organic origin, although not necessarily hereditary: “If an inborn biologic defect exists and plays an important part in such a psychopath's disorder, it is not necessary to assume that the defect is hereditary. Perhaps it may be the result of a subtle failure in maturation, an agenesis of unknown etiology” (p. 412). Cleckley also showed appreciation for the work of Karpman (1941) who maintained that psychogenic factors are responsible for the majority of cases of psychopathy. For a psychoanalyst like Karpman (1948), ‘psychogenic’ largely referred to so-called oedipal conflicts, although he was also attuned to the relevance of parental

maltreatment and rejection in the etiology of ‘secondary psychopathy’, as he termed it. This secondary psychopathy should be distinguished from primary psychopathy which, according to Karpman (1948), was present in only a minority of cases (15%; p. 487) in which an inborn or constitutional defect caused the disorder.

Recently, several scholars have returned to Karpman’s concept of secondary psychopathy (Daversa, 2010) or acquired callousness (Kerig et al., 2012), focusing on the experience of childhood psychological trauma. Earlier, Porter (1996) had already provided an alternate causal pathway to the predominant view of a genetic predisposition to psychopathy of that time. He hypothesized that (secondary) psychopathy is the result of “ ‘de-activation’ or dissociation of a developing basic affective nature and conscience” (p. 183). This de-activation is viewed as a coping mechanism in response to traumatic interpersonal experiences. Porter (1996) linked the affective blunting in psychopathy to the emotional numbing and feelings of detachment that are key diagnostic symptoms of Posttraumatic Stress Disorder (American Psychiatric Association, 2013).

Other theoretical formulations, notably complex trauma (Ford, 2005) and Betrayal Trauma Theory (BTT; Freyd, 1996) are also relevant in relation to the etiology of psychopathy. Complex trauma is defined as exposure to traumatic stressors at an age (e.g., early childhood) or in a context (e.g., prolonged torture or captivity) that compromises secure attachment with primary caregivers (Cook et al., 2005). Complex trauma causes the organism to enter into “survival mode”, resulting in changes to many developing self-regulatory mechanisms (Ford, 2005), including: (a) attention and learning; (b) sensorimotor functions; (c) working (short-term processing), declarative (verbal) and narrative (autobiographical) memory; and (d) emotion regulation and social relatedness (attachment). BTT emphasizes the interpersonal context in

which most early childhood abuse takes place, particularly intrafamilial abuse. Central to BTT is attachment theory, the notion that humans depend on others for physical survival and emotional responsiveness (Bowlby, 1969). According to BTT, the degree to which the abuse represents a betrayal by the trusted and needed attachment figure mediates the way in which abuse-related information is processed and remembered (Sivers et al., 2002). Forgetting and misremembering of the abuse, as well as emotional dissociation, are hypothesized to be specifically linked to betrayal vs. other, nonbetrayal trauma (DePrince et al., 2012).

In summary, there appears to be a recent upsurge in theoretical models that assign an etiological role to early childhood trauma in the development of psychopathy. Several theorists proposed models in which early maltreatment experiences, particularly those with primary caregivers, produce a blunted or dissociative response to stress, as a key factor in the affective deficits observed in psychopathy (e.g., Daversa, 2010; Kerig et al., 2012). Childhood trauma has also been linked to Karpman's 'secondary psychopathy', which is embodied by the present-day behavioral disinhibition/lifestyle facet of psychopathy (Ford et al., 2012).

Empirical Studies about the Role of Childhood Maltreatment in Psychopathy

Research on the relationship between childhood maltreatment and psychopathology, including psychopathy, is mired by a number of methodological difficulties. Ideally, child maltreatment would be measured prospectively and objectively, before the onset of psychopathy. However, in studies of adults, childhood maltreatment is mostly measured through retrospective self-report, which can be biased in the form of memory errors of both omission and commission (Jaffee, 2017; Widom & Czaja, 2012; Widom & Morris, 1997). Other research designs use official Child Protection Services records, but these are likely to underestimate the true prevalence of maltreatment (Jaffee, 2017). As best-evidence practice, childhood maltreatment

would be established through triangulation, using official records, collateral informants, and self-report, but this type of rigor is still quite rare.

The first studies that explored the hypothesized link between childhood maltreatment and psychopathy were case studies conducted in the 1940s (e.g., Bowlby, 1944; Karpman, 1948). Weiler and Widom (1996) performed one of the first prospective studies that used the PCL-R in a large sample of young adults ($n = 652$) and compared them with a control group, matched for demographics and criminal history ($n = 489$). The abused and neglected group was composed of substantiated cases of childhood physical and sexual abuse and/or neglect processed in the county juvenile or adult criminal court and the subjects were 11 years or younger at the time of receiving maltreatment. Abused/neglected individuals had significantly higher PCL-R total scores than matched controls, regardless of gender or ethnicity. Lang et al. (2002) attempted to replicate the Weiler and Widom study with a Swedish sample, comprised of 192 boys aged 11-14 years accused of property crimes and 95 boys, matched for age, family type (separated or not), social group and neighborhood. Child abuse and neglect was determined at age 11-14 by means of triangulation, using school and social work records, parent and child interviews and self-report. The sample was followed up at age 32-40. Mean PCL-R scores of the high victimization group were significantly higher than those of the low victimization group. Notably, neither of these studies examined associations between the four PCL-R facets and childhood maltreatment.

Another early study into the association between childhood maltreatment and PCL psychopathy used a purely retrospective design in a Scottish prison sample (Marshall & Cooke, 1999). Fifty psychopaths (Mean PCL-R = 29) and 55 nonpsychopaths (Mean PCL-R = 13) were administered the Childhood Experience of Care and Abuse interview (CECA; Bifulco et al.,

1994). The CECA asks about specific experiences and events occurring in childhood as opposed to subjective feelings. The authors found significantly higher scores in the psychopathic group compared to the nonpsychopathic group on parental discipline, parental antipathy, parental neglect, parental control, and psychological abuse. No differences were found for physical and sexual abuse. Using multiple regression analysis, Marshall and Cooke (1999) found that victimization within the family significantly predicted PCL-R Factor 1 scores, whereas societal adversity (school experience, school performance, institutional stay) was the main predictor of PCL-R Factor 2 scores. This was one of the first studies that showed a specific effect for childhood family trauma on the Affective-Interpersonal factor. However, a study that used a relatively similar design (a sample of 615 male American offenders, retrospective reports of child abuse/neglect and PCL-R rated psychopathy) did not replicate this finding (Poythress et al., 2006). Using the Cooke and Michie (2001) 13-item model, these authors found a significant association of childhood maltreatment with the behavioral lifestyle factor, but not with the affective and the interpersonal factors (Poythress et al., 2006).

Gender differences

Psychopathy manifests itself differently in women compared to men (Verona & Vitale, 2018). High psychopathy scores in women correlate with more emotional dysregulation as well as manipulative and sexualized behaviors (Forouzan & Nicholls, 2015; Kreis & Cooke, 2011). The number of studies on the link between childhood abuse and psychopathy in female subjects is limited and findings are currently inconclusive (for an overview, see Verona & Vitale, 2018). Whether childhood maltreatment is similarly related to female and male psychopathy is an open question.

The Present Study

Given the current theoretical models and prior empirical research on the link between childhood maltreatment and psychopathy, we deemed it appropriate to conduct a systematic review and meta-analysis of this literature. We chose to include both Hare PCL-derived scales and self-report instruments, which have been used most prominently in the field. We believed this would on the one hand cast the widest net in terms of allowing us to include a large number of studies, and on the other hand give us the possibility to gain insight into the link between the four facets of psychopathy and childhood maltreatment. We opted to include all types of child maltreatment, but did not include other indicators of childhood adversity, such as socioeconomic adversity or community violence, although we acknowledge that these may act as risk factors for antisocial and psychopathic behavior (Farrington & Bergström, 2018).

By conducting a meta-analysis of the existing research base, we aimed to find answers to the following questions:

1. Is there a significant association between any type of childhood maltreatment and psychopathy in general and as defined by its four facets?
2. Do different types of childhood maltreatment (neglect vs. physical vs. sexual vs. emotional abuse) have similar associations with psychopathy?

We were also interested in conducting several moderator analyses:

3. Does gender (male vs. female) impact the link between psychopathy and childhood maltreatment?
4. Does the psychopathy measure used (PCL-scale vs. other psychopathy measure) impact the link between psychopathy and childhood maltreatment?

5. Does the type of publication (peer reviewed journal articles vs. grey literature) impact the link between psychopathy and childhood maltreatment?
6. Does study design (prospective vs. retrospective) impact the link between psychopathy and childhood maltreatment?
7. Does sample type (clinical/correctional vs. community) impact the link between psychopathy and childhood maltreatment?

Method

Protocol and Open Data

This meta-analysis was pre-registered on March 26, 2018, under the PROSPERO platform, an international prospective register of systematic reviews. The registration can be found on their webpage (<https://www.crd.york.ac.uk/PROSPERO/>) with the ID: CRD42018091678. Furthermore, the raw data are accessible under the Open Science Framework (OSF; <https://osf.io/2kqcw/>).

Inclusion Criteria

Our variables of interest were psychopathy, measured by means of self-report or clinical judgment, and childhood maltreatment as defined by the World Health Organization (WHO):

Child maltreatment is the abuse and neglect that occurs to children under 18 years of age. It includes all types of physical and/or emotional ill-treatment, sexual abuse, neglect, negligence and commercial or other exploitation, which results in actual or potential harm to the child's health, survival, development or dignity in the context of a relationship of responsibility, trust or power. Exposure to intimate partner violence is also sometimes included as a form of child maltreatment. (WHO, 2020, first paragraph)

The following criteria were used to include studies: (1) peer-reviewed articles, grey literature, book chapters/books and dissertations; (2) prospective and retrospective studies; (3) studies reported in the English language. Studies needed to include: (1) community, clinical, or correctional samples of children or adults (male and female) in which both childhood maltreatment and psychopathy were measured; (2) childhood maltreatment: physical abuse, emotional abuse, sexual abuse, and neglect, measured by means of self-report and/or official reports; (3) psychopathy measured by means of self-report or clinician rated scales.

Exclusion Criteria

We excluded: (1) single-case studies, conference abstracts without primary study data, book reviews; (2) papers not written in English; (3) studies which focused on childhood adversities outside the family or primary caretaking environment, such as war trauma or community violence; (4) studies on other types of childhood adversity, such as loss of a parent or serious accidental injury, or traumatic experiences after age 18; and (5) studies without an objective measure of psychopathy.

Unlike initially specified in our PROSPERO registration, studies that focused solely on callous-unemotional (CU) traits were excluded from the current review. This decision was made as the number of located articles investigating psychopathy exceeded our expectation so that the inclusion of CU traits would only add heterogeneity to the synthesis of effect sizes.

Literature Search

Consistent with the Preferred Reporting Items for Systematic Review and Meta-Analysis guidelines (PRISMA; Moher et al., 2009), a systematic literature search of papers published between January 1990 and January 2021 was conducted. The year 1990 was chosen because the

first validation study of Hare's revised version of the Psychopathy Checklist (PCL-R; Hare et al., 1990) was published in 1990. The first search was conducted in February 2018 on PsycINFO, PubMed, and Web of Science and subsequently updated in January 2021. The following keywords were used: (*trauma** OR *complex trauma** OR *childhood trauma** OR *abus** OR *adverse childhood experience+* OR *neglect* OR *maltreatment* OR *betrayal trauma*) AND (*psychopath+* OR *psychopathic* OR *psychopathy* OR *callous unemotional trait+* OR *CU trait+* OR *CU-trait+* OR *ODD* OR *CD* OR *conduct disorder+* OR *oppositional defiant disorder+* OR *PCL** OR *antisocial* OR *dissocial*).

To identify all potential and ongoing research, the project was uploaded to ResearchGate (<https://www.researchgate.net/project/The-link-between-childhood-trauma-and-psychopathy-A-systematic-review>). Additionally, researchers who had previously conducted research on psychopathy were contacted via email. Covidence software (www.covidence.org) was utilized to aid the process of the systematic review management and collaboration.

Study Selection

The screening process was conducted in two stages. At each stage, two members of the research team independently assessed whether a publication should be included or excluded for further analysis. A consensus was reached if both assessors agreed on the inclusion or exclusion of the publication. Conflicts were resolved by the first author (CdR). Initially, studies were screened for eligibility by examining their title and abstract. Subsequently, eligible studies were assessed via full text screening. A number of dissertations and grey literature that could not be accessed from open sources were retrieved by contacting the University library in which the dissertation was defended or by contacting the authors through LinkedIn or other social media.

Data Collection

The following information was retrieved from all studies that met the aforementioned inclusion criteria: type of publication; sample type; sample mean age and standard deviation; sample sex distribution; country of study; total sample size; childhood maltreatment measure; and psychopathy measure. On occasion, not all of the above data could be obtained from the text. In such instances, the corresponding author of the paper was contacted for necessary information. Data extraction per study was always carried out by a pair of authors who had to be in agreement. Effect size extraction was conducted by the second author (MB).

Summary Measure

The Pearson product-moment correlation coefficient r was chosen as the effect size index because most studies reported correlational data to quantify the association between childhood maltreatment and psychopathy. If possible, coefficients were extracted directly from zero-order correlation matrices. Missing effect sizes were acquired either by contacting the corresponding authors or by estimating their magnitude from statistical information provided in the paper, such as t -statistics or d values (Lipsey & Wilson, 2001). For three studies (Fisher, 2003; Krstic et al., 2016; Young & Widom, 2014), correlation coefficients had to be imputed from the standardized regression coefficients (Peterson & Brown, 2005).

Given that each paper could report more than one effect size, the following approach was used to ensure independence among effect sizes (Borenstein et al., 2011): (1) independent samples within the same publication were treated as separate studies; (2) in cases where an author published multiple manuscripts, but used samples that were clearly drawn from the same population, only the study with the largest sample was included; (3) if applicable, separate effect sizes for general maltreatment, physical abuse, emotional abuse, sexual abuse, and physical and

emotional neglect were derived from the same paper; (4) similarly, effect sizes for psychopathy total scores as well as for psychopathy facets were extracted; and (5) composite effect sizes for psychopathy were calculated following the recommendations of Borenstein et al. (2011) when multiple psychopathy measures were used or when estimates were only reported for psychopathy factors.¹

In an attempt to address the attenuation of correlation coefficients caused by measurement error (Schmidt & Hunter, 2014), additional corrected effect sizes and variances were computed. The required reliability estimates were retrieved from the included articles, or if the information was lacking, they were imputed using the *Mice* package in R by averaging the results of five imputed data sets (Groothuis-Oudshoorn & Van Buuren, 2011).

Meta-Analysis

Because the studies were expected to vary fundamentally, random-effects models were chosen to calculate the average summary effect sizes (Borenstein et al., 2011). The between-study variance τ^2 was estimated with restricted maximum likelihood method (REML) and used to assign weights to each study by the inverse of total variance: $w_i = \frac{1}{(v_i + \tau^2)}$. Simulations have shown that REML tends to be less biased than other popular methods (Veroniki et al., 2016). While τ^2 reflects the variance of true effects in absolute terms, I^2 was examined to quantify the relative amount of true heterogeneity among the total variability across studies. By convention, I^2 values of 25%, 50%, and 75% respectively indicate low, moderate, and high levels of inconsistency in a meta-analysis (Higgins et al., 2003).

¹ A sensitivity analysis was performed to test whether psychopathy composite effect sizes differed significantly from effect sizes generated by psychopathy total scores. No significant differences were found.

Five separate meta-analyses were performed, for general maltreatment, physical abuse, emotional abuse, sexual abuse, and neglect. Moreover, a priori specified moderator analyses were conducted to determine whether the association between childhood maltreatment and psychopathy differed by (1) proportion of females in the sample; (2) type of sample (clinical/correctional vs. community sample); (3) psychopathy measure (PCL-scale vs. other psychopathy measure); and (4) type of publication (journal article vs. grey literature). Mixed-effect meta-regression models were fitted with categorical moderators included as dummy variables and the proportion of females as a continuous variable (ranging from 0 to 100%).

Given the multifaceted structure of psychopathy (Neumann et al., 2005), additional subgroup analyses were carried out to examine whether the summary effect sizes vary across psychopathy facets (i.e., affective, interpersonal, lifestyle, and antisocial). That is, each facet was included as a subgroup and tested for between-group differences ($H_0: r_{\text{affective}} = r_{\text{interpersonal}} = r_{\text{lifestyle}} = r_{\text{antisocial}}$; Rubio-Aparicio et al., 2020). Significant results were then further investigated with post-hoc analyses contrasting $r_{\text{affective}}$ VS. $r_{\text{lifestyle}}$; $r_{\text{affective}}$ VS. $r_{\text{antisocial}}$; $r_{\text{interpersonal}}$ VS. $r_{\text{lifestyle}}$; $r_{\text{interpersonal}}$ VS. $r_{\text{antisocial}}$. However, since most studies reported effect sizes for all four facets, three-level random effects models were used. Three-level meta-analyses allow for dependencies among effect sizes by breaking τ^2 down into the variance within samples (σ^2_{level2}) and the variance between samples (σ^2_{level3} ; Cheung, 2014).

To test the robustness of the results, included studies were assessed for influential effect sizes by the inspection of multiple influence measures (DIFFITS, Cook's distance, covariance ratio) generated by the leave-one-out method (Viechtbauer & Cheung, 2010). Subsequent sensitivity analyses were performed to determine whether the deletion of such influential cases would alter the overall findings of this meta-analysis. If not, the results can be considered robust

(Viechtbauer & Cheung, 2010). All analyses were performed with the *metafor* package (Viechtbauer, 2010) in the latest version of R 4.0.3 (R Core Team, 2020).

Publication Bias

As with any meta-analysis, publication bias poses a risk to the interpretability of the pooled effect size (Rothstein et al., 2005). Therefore, its presence and impact were examined in three ways. First, funnel plots were created and inspected visually for asymmetry. An asymmetrical distribution of studies across the mean effect size may suggest publication bias (Sterne et al., 2011). Second, Egger's linear regression method was used to test for a linear relationship between the effect size and its standard error, which provides a less subjective measure of asymmetry in a funnel plot (Sterne et al., 2005). Third, Duval and Tweedie's Trim and Fill method was applied to estimate an unbiased summary effect size by imputing missing effect sizes (Duval & Tweedie, 2000). This approach allows one to quantify the impact of reporting bias on the observed result (Duval, 2005).

Results

Descriptive Results

Figure 1 provides a visual outline of the study selection process. Our literature search generated 28,368 papers. Subsequently, 5,920 duplicates were removed. The remaining 22,448 papers entered the title and abstract screening stage, in which 22,208 were excluded in accordance with the inclusion and exclusion criteria specified in the Method section. A total of 240 papers entered the full text screening stage, in which 193 were excluded. The reasons for exclusion were the following: The association between primary variables of interest was not explored ($n = 124$), full text not available ($n = 31$), not a primary study ($n = 23$), contains duplicate data ($n = 9$), insufficient statistical information to calculate effect size ($n = 4$), full text

not in English ($n = 1$), and single case study ($n = 1$). The final sample included in the meta-analysis comprised 55 independent samples and a total of 349 effect sizes originating from 47 papers.

Study and Sample Characteristics

All papers were published between 1996 and 2020. Forty-three of the included papers were journal articles and four were dissertations. The included papers originated from six continents, namely North America ($n = 32$), Europe ($n = 11$), South America ($n = 1$), Africa ($n = 1$), Asia ($n = 1$), and Australia ($n = 1$).

A total of 12,737 participants were included in the final set of articles, where sample sizes ranged from relatively small ($n = 22$; Craparo et al., 2013) to very large ($n = 1,169$; Watts et al., 2017). The included papers predominantly used correctional or clinical samples ($n = 34$), though some papers used samples from the general population ($n = 13$). Twenty-one study samples were solely male, five were female, and 21 were mixed. Further details on study and sample characteristics are given in Table A.1 in the Appendix.

Measure Characteristics

The included papers used a range of measures to explore the variables of interest, namely childhood maltreatment and psychopathy. To assess childhood maltreatment, most papers used retrospective self-report measures, such as the Childhood Trauma Questionnaire (CTQ; Bernstein et al., 1997). A minority of papers used official child welfare records as a measure of childhood maltreatment. Few papers used a combination of self-report measures and official records. For the assessment of psychopathy, most papers employed a clinician-rated scale from the family of Hare Psychopathy Checklist measures, such as the PCL-R (Hare, 1991, 2003), PCL:SV (Hart et al., 1995), or PCL:YV (Forth et al., 2003). A minority of papers used self-

report psychopathy scales, such as the Self-Report Psychopathy Scale (e.g., SRP; Paulhus et al., 2009) or the Youth Psychopathy Traits Inventory (YPI; Andershed et al., 2007). None of the included papers used a combination of clinician-rated and self-report measures to assess psychopathy.

Meta-Analytical Results

The pooled effect sizes for general maltreatment, physical abuse, emotional abuse, sexual abuse, and neglect are presented in Table 1. In what follows, the findings for each type of childhood maltreatment are discussed separately. Corresponding figures, such as forest plots (Figure S.1-S.5) and funnel plots (Figure S.6) can be found in our supplementary material.

<<Insert Table 1 about here>>

General Maltreatment

A statistically significant positive correlation between general childhood maltreatment and psychopathy was found ($r = .20$, 95% CI [.16, .24], $p < .0001$). The amount of true heterogeneity among the 32 effect sizes was moderate ($I^2 = 66.95\%$), and one influential effect size was discovered (Borja & Ostrosky, 2013). By excluding this study, the correlation increased to $r = .21$ (95% CI [.17, .24], $p < .0001$) and I^2 decreased to 57.47%.

Physical Abuse

For physical abuse, 32 independent correlation coefficients were synthesized, which produced a statistically significant positive effect size ($r = .19$, 95% CI [.16, .22], $p < .0001$). The level of heterogeneity was low to moderate ($I^2 = 44.14\%$), and one influential correlation was determined (Boduszek et al., 2019). While the level of heterogeneity dropped to a low level $I^2 = 17.83\%$ without the influential case, the pooled effect size increased slightly ($r = .20$, 95% CI [.18, .23], $p < .0001$).

Emotional Abuse

The association between emotional abuse and psychopathy was positive and statistically significant, based on 25 effect sizes, with $r = .15$ (95% CI [.10, .20], $p < .0001$). The heterogeneity among studies was substantial ($I^2 = 77.34\%$). One influential study was detected (Krstic et al., 2016). Removing Krstic et al. (2016) from the model resulted in a larger summary effect size ($r = .17$, 95% CI [.13, .21], $p < .0001$) and considerably reduced the level of heterogeneity to $I^2 = 56.70\%$.

Sexual Abuse

On the basis of 32 effect sizes, a significant but small positive correlation ($r = .10$, 95% CI [.06, .14], $p < .0001$) between sexual abuse and psychopathy was found. I^2 was moderate to large ($I^2 = 65.79\%$), and no influential effect sizes were found.

Neglect

For neglect, 20 effect sizes were combined. This produced a statistically significant positive correlation ($r = .20$, 95% CI [.16, .25], $p < .0001$), and a moderate amount of heterogeneity ($I^2 = 69.47\%$). The leave-one-out method determined one influential study (Cima et al., 2008). After removing it, the pooled effect size increased to $r = .22$ (95% CI [.18, .26], $p < .0001$), and I^2 decreased somewhat ($I^2 = 58.11\%$).

Correction for Attenuation

Fitting the meta-analytical models with correlation coefficients corrected for attenuation produced only negligibly larger pooled effect sizes, with $r_c = .23$ (95% CI [.19, .28]), $r_c = .22$ (95% CI [.18, .26]), $r_c = .17$ (95% CI [.11, .23]), $r_c = .11$ (95% CI [.07, .15]), and $r_c = .25$ (95% CI [.18, .31]) for general maltreatment, physical abuse, emotional abuse, sexual abuse, and neglect, respectively.

Moderator Analyses

Contrary to our expectations, the proportion of females in a sample, the type of psychopathy measure used (i.e., PCL-scale vs. other), and the type of publication (i.e., peer-reviewed journal article vs. grey literature) did not significantly moderate the association between childhood maltreatment and psychopathy (Table 2). The type of sample (i.e., clinical/correctional vs. community), on the other hand, revealed evidence for moderation. Clinical and correctional samples showed smaller, but still significant, positive associations between general maltreatment and psychopathy ($Q_{\text{between}} = 6.88$, $df = 1$, $p = .0087$) compared to community samples (Table 2). The moderator analysis for prospective vs. retrospective studies of child maltreatment could not be carried out due to a lack of prospective studies.

<<Insert Table 2 about here>>

Analyses for the Four Facets of Psychopathy

The analyses on the four facets of psychopathy was conducted on a limited set of studies (k between 9 and 11). Three-level random effects models revealed different patterns for the different types of childhood maltreatment (Table 3). General maltreatment, physical abuse, and neglect showed significant associations with all four facets, although of different magnitudes. Emotional abuse and sexual abuse were significantly correlated with some, but not with all of the four psychopathy facets.

Significant between-group differences across facets were found for general maltreatment ($Q_{\text{between}} = 25.57$, $df = 3$, $p < .0001$), physical abuse ($Q_{\text{between}} = 16.92$, $df = 3$, $p = .0007$), and emotional abuse ($Q_{\text{between}} = 13.28$, $df = 3$, $p = .0041$). Subsequent post-hoc analyses indicated a stronger relationship between childhood maltreatment and the lifestyle and antisocial facets compared to the affective and interpersonal facets. These findings are presented in Table 3.

<<Insert Table 3 about here>>

Publication Bias

The visual inspection of funnel plots provided no indication of asymmetry (Figure S.6). This was corroborated by nonsignificant Egger's regression tests for general maltreatment ($z = -.74, p = .46$), physical abuse ($z = -.38, p = .71$), emotional abuse ($z = -.10, p = .92$), sexual abuse ($z = -1.61, p = .11$), and neglect ($z = -1.74, p = .08$). Furthermore, Duval and Tweedie's Trim and Fill method only imputed missing studies ($k = 8$) for sexual abuse on the right side of the funnel plot (i.e., positive correlations; Figure S.6), resulting in a larger unbiased pooled effect size: $r_{\text{unbiased}} = .13$ (95% CI [.09, .18]).

Discussion

This meta-analysis was conducted against the background of past and present-day theories about the role of childhood maltreatment in the etiology of psychopathy. Theorists have offered diverging views on the purported link between child maltreatment and the different components of psychopathy. Some (e.g., Ford et al., 2012; Karpman, 1948) have proposed a link between childhood maltreatment and the behavioral disinhibition/impulsivity aspect of psychopathy, whereas others (e.g., Daversa, 2010; Porter, 1996) also hypothesized an association between the affective component and child maltreatment. As such, the separate components of psychopathy are viewed as the result of distinct symptomatic responses to psychological trauma, with psychopathic behavioral disinhibition linked to increased arousal and reactivity (cf. DSM-5 criterion E for PTSD; American Psychiatric Association, 2013), and psychopathic lack of affect linked to emotional numbing/avoidance/dissociation (DSM-5 criterion C for PTSD; American Psychiatric Association, 2013).

The association between general childhood maltreatment and psychopathy was $r = .20$ ($r_c = .23$). For the specific types of child maltreatment, the associations were $r = .19$ ($r_c = .22$) for physical abuse, $r = .15$ ($r_c = .17$) for emotional abuse, $r = .10$ ($r_c = .11$) for sexual abuse, and $r = .20$ ($r_c = .25$) for neglect. There has been discussion in the field on how to interpret effect size r . Cohen (1988) recommended Pearson r values of .10, .30, and .50 to demarcate small, medium, and large effects, respectively. However, recent research has cast doubt on this classification. A review of studies by Gignac and Szodorai (2016), based on 708 meta-analytically derived correlations, reported that the 25th, 50th, and 75th percentiles corresponded to correlations of .11, .19, and .29, respectively. Fewer than 3% of correlations met Cohen's definition of 'large' (i.e., .50 or higher). Gignac and Szodorai (2016) suggest that in real life, the terms small, medium and large more closely correspond to correlations of .10, .20 and .30. If we take the latter values as benchmarks, psychopathy has a medium sized association with childhood maltreatment in general, physical abuse, and neglect, and a small association with childhood emotional abuse and sexual abuse.

Eleven studies in the meta-analysis included data on the four facets of psychopathy in relation to childhood maltreatment, and a significant variation was revealed across all four facets for general maltreatment, physical abuse, and emotional abuse. Overall, the associations for the lifestyle and antisocial facets were stronger than for the affective and interpersonal facets, with the exception of neglect and sexual abuse. Of note, sexual abuse only showed a significant but small correlation with the lifestyle facet. The summary effect sizes of general childhood maltreatment, physical abuse, and neglect were generally larger, and mostly in the moderate range for the lifestyle and antisocial facet, and in the small range for the affective and interpersonal facet. As such, these findings support theoretical notions that both the

affective/interpersonal and the behavioral/antisocial components of psychopathy are related to experiences of childhood maltreatment in primary attachment relationships (Daversa, 2010; Kerig et al., 2012; Larstone et al., 2018), but likely to a different degree. It should be noted that these facet analyses were conducted on a much smaller sample of effect sizes, which limits the robustness of these specific findings. Conclusions regarding the link between different types of childhood maltreatment and the four facets of psychopathy should be deferred until further studies become available.

Our moderator analysis showed, contrary to our expectations, that gender of study participants, type of psychopathy measure (PCL-scale vs. other) used and type of publication (peer-reviewed vs. grey literature) did not impact the association between psychopathy and childhood maltreatment. As such, a reasonable inference is that the childhood maltreatment-psychopathy relationship is relatively robust across levels of these moderators, including gender, psychopathy measure, and publication type. We should mention that the number of effect sizes per subgroup (k) for some of these moderator analyses, was rather small, which limits the robustness and interpretability of our findings. We found significant moderation for sample type: correctional and clinical samples showed smaller, but still significant, correlations between psychopathy and general childhood maltreatment compared to general community samples (.16 vs. .26). The interpretation of this finding is not straightforward. On one hand, one could argue that the opposite effect would have been more logical, because psychopathic traits and child maltreatment are less prevalent in community samples than in clinical and correctional samples. However, our results appear in line with the meta-analytic findings of Douglas et al. (2009) who studied the association between psychotic symptoms and violence. They found that

this relationship was significantly stronger in community samples compared to psychiatric and correctional samples.

Our meta-analysis points at significant gaps in the research literature: most studies have been conducted in institutionalized clinical/correctional samples, using retrospective self-report of childhood maltreatment. Therefore, it is possible that memory distortion (so-called recall bias) may impact the strength of the association between psychopathy and childhood maltreatment. Prospective studies continue to be rare, likely because they require extreme efforts and are thus, costly. A relevant example is a longitudinal study by Shi et al. (2012), which found that actual, observed maternal withdrawal in response to an 18-month old infant's distress (cf., emotional neglect), significantly predicted features of Antisocial Personality Disorder (ASPD) 20 years later. Although not identical, ASPD and psychopathy are similar, and it is interesting to note how a study with a longitudinal design that used such a long follow-up period provides support in line with our hypothesis.

An important issue in meta-analytic research is whether the reported findings are robust and valid. The number of studies included in the present meta-analysis is reasonably high (between 20 and 32). We also succeeded in including studies from different continents, although North-American studies dominated. Both the number of included studies and their global representativeness increase our trust in the robustness of the effects. In addition, the robustness of our findings was further supported by our sensitivity analysis suggesting that our results are not substantially impacted by a few influential cases. A common threat to the validity of meta-analytic findings is the file-drawer problem. Published research studies may overestimate the true effect sizes if journals prefer to accept papers that report strong significant associations over papers with nonsignificant or small effects (Lipsey & Wilson, 2001). Comparisons of the results

of published and unpublished studies as well as the trim and fill procedure (Duval & Tweedie, 2000) and Egger's regression test for funnel plot asymmetry (Sterne et al., 2005) indicated that the file drawer issue was not a major concern in our meta-analysis.

Implications for Theory

The present meta-analysis provides support for a moderate link between psychopathy and childhood physical abuse, emotional abuse, and neglect, as well as overall childhood maltreatment. The link between psychopathy and childhood sexual abuse is small. These associations are stronger for the lifestyle and antisocial facets than for the affective and interpersonal facets of psychopathy, but nearly all associations are significant. Our findings are in line with theories of the impact of complex trauma (Ford, 2005) and BTT (Freyd, 1996) on the development of serious personality pathology, although most previous theorizing and empirical research have focused on Borderline Personality Disorder (for a review, see Luyten et al., 2020) and not psychopathy. An exception to this is a recent study by Yalch and Levendosky (2020) who found that exposure to trauma high in betrayal was the only predictor of the vulnerable and grandiose dimensions of pathological narcissism in a college student sample, after controlling for other forms of (nonbetrayal) trauma exposure. The authors suggest "that not only does exposure to high betrayal trauma inflict a psychological wound (narcissistic vulnerability), but also that it influences the means by which people defend against that wound (narcissistic grandiosity)" (Yalch & Levendosky, 2020, p. 1041). Narcissistic grandiosity resembles the interpersonal facet of psychopathy.

The view that at least some of the signature features of psychopathy can be seen as responses to complex or betrayal trauma, that is, repeated incidents of maltreatment over an extended period of time (i.e., months or years) which includes emotional abuse, physical abuse,

neglect, and/or witnessing family violence within the caregiver system, provides further direction for preventive and therapeutic efforts. Therapeutic interventions that focus on early childhood trauma can provide a test of the causal role of trauma, if it could be shown that processing of traumatic experiences leads to meaningful reductions in psychopathic traits.

Several trauma-focused treatment interventions for psychopathy have been developed: Dialectical Behavior Therapy for psychopathy (Galietta & Rosenfeld, 2012) and Schema Therapy (ST) for forensic patients with personality disorders, including psychopathy (Bernstein et al., 2007). Controlled effectiveness studies of these therapy models have not yet been published, but a single case study documented the process of individual ST in a Dutch forensic patient with psychopathic traits (Chakhssi et al., 2014). This patient had been a victim of extreme physical and emotional abuse as a child and the therapist used different ST techniques (e.g., limited reparenting, experiential techniques) to alter the patient's maladaptive schema modes. The case study also showed the patient's PCL-R total score changed from 27 at baseline to 14 after four years of intensive ST. Remarkably, the Affective facet showed the largest change: from 7 to 1; the Interpersonal facet decreased from 4 to 1. This finding, although just an $N = 1$ result, challenges the notion that affective and interpersonal features of psychopathy are immutable (Olver, 2016; see de Ruiter & Hildebrand, in press for further argumentation).

Limitations

The findings of this systematic and meta-analytic review should be considered in light of a number of limitations. First and foremost, correlational analyses, such as those conducted in this meta-analysis, cannot be used to prove causality. Shadish, Cook, and Campbell (2002) summarized John Stuart Mill's three criteria for inferring causality: "A causal relationship exists if (1) the cause preceded the effect, (2) the cause was related to the effect, and (3) we can find no

plausible alternative explanation for the effect other than the cause” (p. 6). Thus, the clearest evidence for a causal relationship comes from experimental rather than correlational research. From a strictly methodological standpoint, none of the studies included in our meta-analysis, not even those that used prospective measures of childhood maltreatment, such as child protection data, fulfill all of the three criteria. Potentially confounding variables were not included in most studies. A majority of the studies used a cross-sectional design in which psychopathic offenders retrospectively reported more childhood abuse than nonpsychopathic offenders. We cannot rule out the alternative hypothesis, that psychopathy caused childhood maltreatment because there is a possibility that psychopathic traits may have caused these individuals to retrospectively report higher rates of childhood maltreatment. The latter effect is generally referred to as ‘recall bias’ (Widom et al., 2004). It occurs when the accuracy and inaccuracy in reporting prior experiences vary as a function of present (physical or psychological) health condition. “One process believed to underlie the differential reporting accuracy is ‘effort after meaning’, where unhealthy individuals exert more effort to search for disease explanation and assign more meaning to past events” (Widom et al., 2004, p. 718). Thus, people with psychopathic traits could attribute their problems in life to childhood maltreatment.

Second, the operationalization of childhood maltreatment varied greatly from study to study. In addition to different assessment measures, researchers also used diverse definitions of different types of child maltreatment. It would greatly benefit this field of study if scholars would agree on the operational definitions of the different types of childhood maltreatment. Perhaps, the definition of child maltreatment by the World Health Organization (Meinck et al., 2016) could serve as a starting point. The 1999 WHO Consultation Group on Child Abuse Prevention distinguish four types of child maltreatment: physical abuse, emotional (or psychological) abuse,

neglect, and sexual abuse. Recently, exposure to domestic violence has also been recognized as a separate form of child maltreatment by the WHO (Meinck et al., 2016). Additionally, the context of child maltreatment should be clearly operationalized in future studies because present day theories (e.g., Ford, 2005) clearly distinguish between the psychological consequences of (chronic) maltreatment by primary caregivers versus other maltreatment of children (e.g., a single incident of sexual abuse by a stranger).

Third, although we managed to retrieve studies from six continents, samples from Western countries were clearly overrepresented in our meta-analysis. This has very likely impacted the findings, because the two variables we studied, psychopathy and childhood maltreatment, are not impervious to ethnic and cultural factors. A recent review of cultural and ethnic variations in psychopathy (Fanti et al., 2018) concluded that there is evidence for ethnic differences in total psychopathy and facet scores, for instance, between African American and European American prisoners, as well as differential responding in laboratory tasks of emotional and cognitive processing between these groups. Thus, ethnic differences may exist in the way psychopathy manifests itself and its underlying mechanisms, possibly including its etiology. Furthermore, what is considered child maltreatment varies according to socially accepted norms, which are heavily influenced by culture. Certain parenting styles, such as corporal punishment, are viewed as inappropriate in some cultures, but are accepted and even promoted in others (Earner, 2007; Hassan & Rousseau, 2009). Additionally, opinions about what constitutes child abuse have been found to differ amongst cultures (e.g., Boakye, 2009; Lansford et al., 2015; Plummer & Njuguna, 2009). How cultural and ethnic factors shape the relationship between childhood maltreatment and psychopathic traits warrants further study.

Conclusions

The current meta-analysis reports moderate effect sizes between psychopathy and childhood maltreatment. Effects are stronger for the behavioral lifestyle and antisocial facets than for the affective and interpersonal facets. We found that the associations of psychopathic traits with child maltreatment were largely invariant across gender, type of psychopathy measure, and publication type. The sample type was found to be a moderating factor.

Our findings support theoretical models and empirical research that suggest a role of childhood trauma, and childhood maltreatment by primary caregivers in particular, in the etiology of psychopathy. This role may be somewhat larger in the behavioral lifestyle (or emotional dysregulation) symptoms of psychopathy than in the affective (or emotional numbing) and interpersonal (narcissistic) symptoms, but still relevant to both. Future research in this domain needs to focus on prospective, longitudinal designs across extensive time spans, because retrospective, cross-sectional designs cannot inform us about causal directions. Furthermore, a common set of childhood maltreatment measures, including behavioral observations would facilitate cross-study comparisons and future meta-analyses. Finally, experimental studies of the “trauma-psychopathy hypothesis”, such as evaluations of trauma-informed therapeutic interventions with children, adolescents, and adults with psychopathic traits will provide further insight into the mechanisms underlying this complex disorder.

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Tables

Table 1

Meta-analytical results of the association between childhood maltreatment and psychopathy

Type of maltreatment	<i>k</i>	<i>r</i>	SE _{<i>r</i>}	95% CI	<i>I</i> ² (%)
General maltreatment	32 (31)	.20 ^{***} (.21) ^{***}	.02 (.02)	[.16, .24] ([.17, .24])	66.95 (57.47)
Physical abuse	32 (31)	.19 ^{***} (.20) ^{***}	.02 (.01)	[.16, .22] ([.18, .23])	44.14 (17.83)
Emotional abuse	25 (24)	.15 ^{***} (.17) ^{***}	.03 (.02)	[.10, .20] ([.13, .21])	77.34 (56.70)
Sexual abuse	32	.10 ^{***}	.02	[.06, .14]	65.79
Neglect	20 (19)	.20 ^{***} (.22) ^{***}	.02 (.02)	[.16, .25] ([.18, .26])	69.47 (58.11)

Note. Results without influential cases are in parentheses.

k = number of effect sizes; *r* = pooled correlation; SE_{*r*} = standard error of *r*; CI = confidence interval of *r*; *I*² = proportion of true heterogeneity.

*** *p* < .001.

Table 2

Results of mixed effects meta-regression models assessing potential moderators by each type of childhood maltreatment

Moderator	<i>k</i>	<i>r</i>	SE _{<i>r</i>}	95% CI	<i>Q</i>	
					between	within
General maltreatment						
Psychopathy measure					.42	74.08***
PCL	25	.19***	.02	[.15, .24]		
Other	7	.22***	.04	[.15, .29]		
Type of publication					.54	79.35***
Journal article	31	.20***	.02	[.16, .24]		
Grey literature	1	.30*	.14	[.03, .57]		
Type of sample					6.88**	63.29***
Clinical/correctional	23	.17***	.02	[.12, .21]		
Community	9	.26***	.03	[.20, .32]		
Proportion females	32	.07 ^a	.05 ^a	[-.03, .16] ^a	1.79	77.82***
Physical abuse						
Psychopathy measure					1.07	50.20*
PCL	22	.21***	.02	[.16, .25]		
Other	10	.17***	.03	[.12, .22]		
Type of publication					.22	51.11**
Journal article	29	.19***	.02	[.15, .22]		
Grey literature	3	.21***	.05	[.12, .31]		
Type of sample					.06	51.46**
Clinical/correctional	24	.19***	.02	[.15, .23]		
Community	8	.19***	.03	[.13, .24]		
Proportion females	32	-.03 ^a	.04 ^a	[-.10, .05] ^a	.40	49.04*
Emotional abuse						
Psychopathy measure					1.09	94.65***
PCL	15	.12**	.04	[.05, .20]		
Other	10	.18***	.04	[.10, .25]		
Type of publication					.19	104.05***
Journal article	23	.15***	.03	[.09, .20]		
Grey literature	2	.19*	.09	[.01, .36]		

Type of sample					.00	104.78 ^{***}
Clinical/correctional	17	.15 ^{***}	.04	[.08, .22]		
Community	8	.15 ^{**}	.05	[.06, .24]		
Proportion females	25	-.04 ^a	.06 ^a	[-.16, .09] ^a	.36	101.18 ^{***}
Sexual abuse						
Psychopathy measure					.13	90.56 ^{***}
PCL	19	.09 ^{**}	.03	[.03, .14]		
Other	13	.10 ^{***}	.03	[.05, .16]		
Type of publication					.78	88.10 ^{***}
Journal article	29	.04	.06	[-.08, .17]		
Grey literature	3	.10 ^{***}	.02	[.06, .14]		
Type of sample					.83	87.85 ^{***}
Clinical/correctional	22	.08 ^{***}	.03	[.03, .13]		
Community	10	.12 ^{***}	.03	[.06, .18]		
Proportion females	31 ^b	.05 ^a	.05 ^a	[-.05, .14] ^a	.98	90.51 ^{***}
Neglect						
Psychopathy measure					.62	59.04 ^{***}
PCL	13	.22 ^{***}	.03	[.16, .29]		
Other	7	.18 ^{***}	.04	[.11, .25]		
Type of publication					.24	59.15 ^{***}
Journal article	17	.20 ^{***}	.03	[.15, .25]		
Grey literature	3	.23 ^{***}	.07	[.10, .36]		
Type of sample					2.81	47.89 ^{***}
Clinical/correctional	14	.17 ^{***}	.03	[.12, .23]		
Community	6	.25 ^{***}	.03	[.18, .31]		
Proportion females	20	.02 ^a	.06 ^a	[-.11, .15] ^a	.10	56.89 ^{***}

Note. k = number of effect sizes; r = pooled correlation; SE_r = standard error of r ; CI = confidence interval of r ; Q_{between} = variance accounted for by subgroups; Q_{within} = residual heterogeneity in effect size.

^aUnstandardized regression coefficient. ^bOne study was excluded from the model as the proportion of females in this study was not reported.

*** $p < .001$. ** $p < .01$. * $p < .05$

Table 3

Between-group differences among individual psychopathy factors by each type of childhood maltreatment

	<i>k</i>	<i>r</i>	<i>SE_r</i>	95% CI	<i>Q</i> _{between}	<i>Q</i> _{within}	Post-hoc comparisons (<i>z</i> -value)			
							aff vs. lif	aff vs. ant	int vs. lif	int vs. ant
General maltreatment					25.57***	78.20***	2.99**	3.24**	3.81***	4.06***
Affective	11	.11***	.03	[.05, .18]						
Interpersonal	11	.08**	.03	[.02, .15]						
Lifestyle	11	.21***	.03	[.15, .28]						
Antisocial	11	.22***	.03	[.16, .29]						
Physical abuse					16.92***	37.88	1.45	3.86***	.66	3.07**
Affective	11	.09**	.03	[.03, .15]						
Interpersonal	11	.12***	.03	[.06, .18]						
Lifestyle	11	.14***	.03	[.08, .20]						
Antisocial	11	.23***	.03	[.17, .28]						
Emotional abuse					13.28**	61.72**	2.58**	1.86	3.13**	2.40*
Affective	10	.07	.04	[.00, .15]						
Interpersonal	10	.05	.04	[-.02, .13]						
Lifestyle	10	.17***	.04	[.10, .25]						
Antisocial	10	.15***	.04	[.07, .22]						
Sexual abuse					6.60	41.09	–	–	–	–
Affective	11	.00	.03	[-.06, .07]						
Interpersonal	11	.03	.03	[-.03, .10]						
Lifestyle	11	.09**	.03	[.03, .16]						
Antisocial	11	.03	.03	[-.03, .09]						

Neglect					3.98	35.00	–	–	–	–
Affective	9	.12***	.03	[.06, .18]						
Interpersonal	9	.12***	.03	[.06, .18]						
Lifestyle	9	.17***	.03	[.12, .23]						
Antisocial	9	.17***	.03	[.11, .23]						

Note. k = number of effect sizes; r = pooled correlation; SE_r = standard error of r ; CI = confidence interval of r ; Q_{between} = variance accounted for by subgroups; Q_{within} = residual heterogeneity in effect size; aff = affective; int = interpersonal; lif = lifestyle; ant = antisocial.

*** $p < .001$. ** $p < .01$. * $p < .05$.

Figures

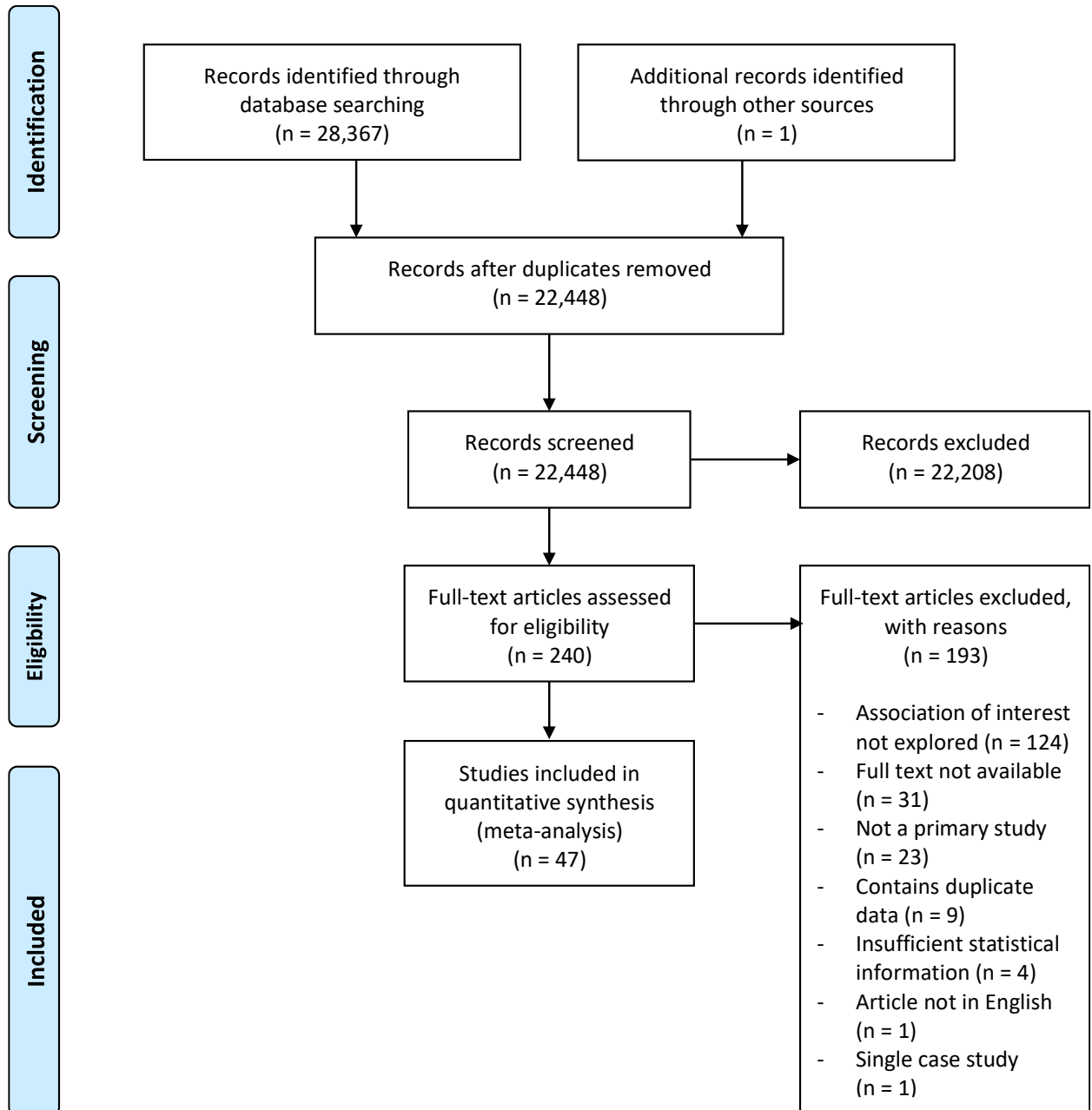


Figure 1. PRISMA flowchart showing the study selection process.

Appendix A

Table A.1

Study characteristics of all included papers

Article	N	% females	Country	Sample type	Publication type	Measure	
						Psychopathy	Childhood maltreatment
Blonigen et al. (2012)	215	100	USA	Correctional	Journal article	PCL-R	PTE
Boduszek et al. (2019)	325	100	Barbados & Grenada	Community	Journal article	PPTS	Self-made questionnaire
Bohle & de Vogel (2017)	418	51	NL	Correctional	Journal article	PCL-R	Questionnaire to assess victimization
Borja & Ostrosky (2013)	194	0	Mexico	Correctional	Journal article	PCL-R	ETI
Campbell et al. (2004)	226	17	Canada	Correctional	Journal article	PCL:YV	AEQ
Cima et al. (2008)	47	0	NL	Correctional	Journal article	PPI	CTQ
Christopher et al. (2007)	142	100	USA	Correctional	Journal article	LSRP	CTQ
Cooke et al. (2020)	789	62	USA	Community	Journal article	LSRP	CMIS-SF/

								Self-made questionnaire
Craparo et al. (2013)	22	0	Italy	Correctional	Journal article	PCL-R		TEC
Dargis et al. (2016)	183	0	USA	Correctional	Journal article	PCL-R		CTQ
Dargis & Koenigs (2018)	222	0	USA	Correctional	Journal article	PCL-R		CTQ
Durand & de Calheiros Velozo (2018)	400	70	USA	Community	Journal article	TriPM		CTQ-SF
Farina et al. (2018); Pennsylvania	253	40	USA	Correctional	Journal article	YPI		CTQ
Fisher (2003)	110	23	USA	Correctional	Dissertation	PSD		Official records
Forouzan & Nicholls (2015)	32	100	Canada	Community	Journal article	PCL-R		Purpose-built protocol
Gao et al. (2010)	333	39	Mauritius	Community	Journal article	SRP-II		CTS
Gao et al. (2011)	71	0	USA	Community	Journal article	PCL-R		CTS
Gowin et al. (2013)	67	15	USA	Clinical/Correctional	Journal article	SRP-III		CTQ

Grady et al. (2019)	105	0	Canada	Correctional	Journal article	PCL-R	Official records/self-report
Graham et al. (2012)	223	0	USA	Correctional	Journal article	PCL-R	Official records
Hong & Lishner (2016)	248	–	USA	Community	Journal article	SRP-III	SLEQ
Jia et al. (2020)	991	76	China	Community	Journal article	DD	CPANS
Kimonis et al. (2012)	373	0	USA	Correctional	Journal article	YPI	LES
Kolla et al. (2014)	24	0	UK	Correctional	Journal article	PCL-R	ETI
Koivisto & Haapasalo (1996)	52	15	Finland	Correctional	Journal article	PCL	Official records
Krischer & Sevecke (2008)	185	48	Germany	Correctional	Journal article	PCL:YV	CTQ
Krstic et al. (2016)	397	0	USA	Correctional	Journal article	PCL-R	MASA
Lang et al. (2002)	199	0	Sweden	Correctional	Journal article	PCL	Official records/self-report
Marshall & Cooke (1999)	105	0	UK	Correctional	Journal article	PCL-R	CECA

McBride (1998); study 1	233	0	Canada	Correctional	Dissertation	PCL:YV	Official records
Moore (2004)	67	0	USA	Correctional	Dissertation	PCL:YV	Official records
O'Neill et al. (2003)	51	0	USA	Clinical/Correctional	Journal article	PCL:YV	CASI
Ometto et al. (2016)	107	44	Brazil	Community	Journal article	PCL:YV	CTQ
Poythress et al. (2006)	615	0	USA	Correctional	Journal article	PCL-R	CATS
Rock (2016)	220	35	USA	Correctional	Dissertation	PCL-SV	FHHQ
Rose et al. (2020)	68	0	USA	Correctional	Journal article	PCL:YV	Self-report
Schimmenti et al. (2015)	78	0	Italy	Correctional	Journal article	PCL-R	TEC
Schraft et al. (2013)	147	14	USA	Correctional	Journal article	PCL:YV	CTQ-SF
Sevecke et al. (2016)	334	50	Germany	Correctional	Journal article	PCL:YV	CTQ
Strand et al. (2016)	80	50	Australia	Correctional	Journal article	PCL:YV	CTQ
Swogger et al. (2012)	75	0	USA	Correctional	Journal article	PCL-R	ACEC

Vahl et al. (2016)	439	0	NL	Correctional	Journal article	YPI	CTQ-SF
Verona et al. (2005)	226	100	USA	Correctional	Journal article	PCL-R	Official records/self-report
Waller et al. (2018)	261	58	USA	Community	Journal article	SRP-SF-IV	CTQ
Watts et al. (2017)	1,169	73	USA	Community	Journal article	PPI-R & LSRP	CTQ
Weiler & Widom (1996)	1,069	50	USA	Community	Journal article	PCL-R	Official records
Young & Widom (2014)	547	58	USA	Community	Journal article	PCL-R	Official records

Note. % females = proportion of females in sample; ACEC = Adverse Childhood Experience Scale; AEQ = Abusive Experience Questionnaire; CASI = Comprehensive Adolescent Severity Index; CATS = Childhood Abuse and Trauma Scale; CECA = Childhood Experience of Care and Abuse; CMI-SF = Childhood Maltreatment Interview Schedule – Short Form; CPANS = Child Psychology Abuse and Neglect Scale; CTQ = Childhood Trauma Questionnaire; CTQ-SF = Childhood Trauma Questionnaire – Short Form; CTS = Conflict Tactics Scale; DD = Dirty Dozen; ETI = Early Trauma Inventory; FHHQ = Family Health History Questionnaire; LES = Life Events Scale; LSRP = Levenson Self-Report Psychopathy Scale; MASA = Multidimensional Assessment of Sex and Aggression; N = total number of participants; PCL-R = Psychopathy Checklist – Revised; PCL-SV = Psychopathy Checklist – Screening Version; PCL:YV = Psychopathy Checklist – Youth Version; PPI = Psychopathy Personality Inventory; PPI-R = Psychopathy Personality

Inventory – Revised; PPTS = Psychopathic Personality Traits Scale; PSD = Psychopathy Screening Device; PTE = Potentially Traumatic Events; SLEQ = Sexual Life Experience Questionnaire; SRP-II/SRP-III = Self-Report Psychopathy Scale – II/III; SRP-SF-IV = Self-Report Psychopathy Scale – Short Form IV; TEC = Traumatic Experience Checklist; TriPM = Triarchic Psychopathy Measure; YPI = Youth Psychopathic Traits Inventory; minus sign (–) = missing data.

Appendix B

References for studies included in the meta-analysis.

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